

# UNIVERSITY OF TWENTE.

## EXAM: Causal Inference

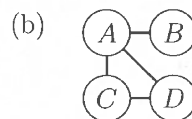
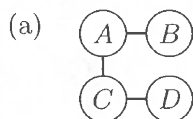
July 1st, 2025, 13.45-16.45 h.

Instructor: Wouter Koolen

Please answer the questions on separate sheets of paper. Do not forget to number the sheets and to write your name on each of them.

### Exercises

- [2pt] Suppose consumption of vitamins has a positive causal effect on life expectancy. Yet in an observational study we see a negative correlation. How can that be explained?
- Consider the following two undirected graphical models.



- [1pt] Does a fully independent joint distribution of  $A, B, C, D$  satisfy the constraints expressed by the undirected graphical model (a)?
  - [1pt] Does any  $P$  that satisfies (b) satisfy (a) as well? What about the converse?
  - [1pt] Give a conditional independence statement that is true under (b).
  - [1pt] Give a conditional independence statement that is true under (a) but does not follow from (b).
- [3pt] Consider observed binary variables  $L, V, A, Y$  and unobserved potential outcomes  $Y^0, Y^1$  satisfying the standard assumptions (P), (C), (CE) recalled on the cheat sheet. We want to estimate the conditional causal effect  $\text{CATE} = \mathbb{E}[Y^1 - Y^0 | V = v]$ . In the lectures we assumed that the group label  $V$  is a coarsening of the covariate  $L$ . Here instead we assume that  $Y^a \perp\!\!\!\perp V | L$  for both treatments  $a \in \{0, 1\}$ . Show that CATE is identifiable.
  - Consider four binary variables: treatment  $A$ , outcome  $Y$  and potential outcomes  $Y^0$  and  $Y^1$  (we ignore covariates in this question). We assume that treatment  $A$  and outcome  $Y$  are observed, while the potential outcomes  $Y^0$  and  $Y^1$  are not. Let us impose the following constraints on the joint distribution  $P(A, Y, Y^0, Y^1)$ :

- (C) Consistency:  $Y = Y^A$   
 (CE) Conditional exchangeability:  $Y^a \perp\!\!\!\perp A$  for each  $a \in \{0, 1\}$   
 (SI) Simple independence:  $Y^0 \perp\!\!\!\perp Y^1$

- (a) [1pt] Do these constraints imply  $Y \perp\!\!\!\perp A$ ? Prove or give a counterexample.  
 (b) [1pt] Show that the joint distribution  $P(A, Y, Y^0, Y^1)$  is uniquely determined by its three marginal probabilities  $\alpha = P(A = 1)$  and  $\gamma^a = P(Y^a = 1)$  for  $a \in \{0, 1\}$ .  
 (c) [1pt] With the same marginal notation as above, show that the marginal distribution  $P(A, Y)$  of the observed variables  $A$  and  $Y$  is

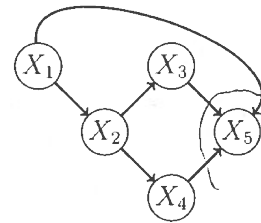
	$Y = 1$	$Y = 0$
$A = 1$	$\alpha\gamma^1$	$\alpha(1 - \gamma^1)$
$A = 0$	$(1 - \alpha)\gamma^0$	$(1 - \alpha)(1 - \gamma^0)$

- (d) [1pt] Can any candidate joint distribution  $Q(A, Y)$  of observed variables  $A, Y$  arise from a distribution  $P$  of the above form?  
 (e) [2pt] Consider the two graphical models below. For each, determine if it is a faithful representation of the constraints on  $P$ .



5. Consider the Linear Structural Equation Model (left) and corresponding DAG (right)

$$\begin{aligned}
 X_1 &= \epsilon_1 \\
 X_2 &= \beta_{21}X_1 + \epsilon_2 \\
 X_3 &= \beta_{32}X_2 + \epsilon_3 \\
 X_4 &= \beta_{42}X_2 + \epsilon_4 \\
 X_5 &= \beta_{51}X_1 + \beta_{53}X_3 + \beta_{54}X_4 + \epsilon_5
 \end{aligned}$$



where  $\epsilon_2, \dots, \epsilon_5$  are i.i.d. standard normal.

- (a) [2pt] Give the Markov equivalence class and the corresponding CPDAG.  
 (b) [1pt] Compute the causal effect per unit intervention  $\frac{\partial}{\partial x} \mathbb{E}[X_5 | do(X_1 = x)]$ .  
 (c) [1pt] Is  $\{X_1, X_3\}$  an adjustment set (aka set of control variables) for the effect of  $X_4$  on  $X_5$ ?  
 (d) [1pt] Assume that  $\beta_{21} \neq 0$ . Show that  $X_1$  is an instrumental variable for the effect of  $X_2$  on  $X_4$ . Do so by proving that  $\text{Cov}(X_4 - \beta_{42}X_2, X_1) = 0$  yet  $\text{Cov}(X_2, X_1) \neq 0$ .